

BLIND SIGNAL PROCESSING FOR TIME-VARYING CONVOLUTIVE MIXING SYSTEMS BASED ON SEQUENCE ESTIMATION ON PARTLY SMOOTH MANIFOLDS

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ABSTRACT

In this paper we focus on Bayesian blind and semi-blind adaptive signal processing based on a broadband MIMO FIR model (e.g., for blind source separation (BSS) and blind system identification (BSI)). Specifically, we study in this paper a framework allowing us to systematically incorporate various types of prior knowledge: (1) source signal statistics, (2) deterministic knowledge on the mixing system, and (3) stochastic knowledge on the mixing system. In order to exploit all possible types of source signal statistics (1), our considerations are based on TRINICON, a previously introduced generic framework for broadband blind (and semi-blind) adaptive MIMO signal processing. The motivation for this paper is threefold: (a) the extension of TRINICON to Bayesian point estimation to address (3) in addition to (1), and (b) more specifically to unify system-based blind adaptive MIMO signal processing with the tracking of time-varying scenarios, and finally (c) to show how the Bayesian TRINICON-based tracking can be formulated as a sequence estimation approach on arbitrary partly smooth manifolds. As we will see in this paper, the Bayesian approach to incorporate stochastic priors and the manifold learning approach to exploit deterministic system knowledge (2) complement one another very efficiently in the context of TRINICON.

Index Terms— Bayesian learning, tracking, manifolds, convolutive BSS, blind system identification

1. INTRODUCTION

The analysis of scenes using sensor arrays (e.g., based on multiple channels of acoustic signals, sonar signals, or radar signals) is a fundamental problem in signal processing with many potential applications and use cases, e.g., for localization/object tracking, source separation/extraction, characterization of the environment, and signal enhancement/dereverberation. In general, the received sensor signals may consist of a mixture of components from multiple simultaneously active sources or objects of interest and various interfering signals located at different spatial positions. Additionally, there can be multipath and/or dispersive propagation of the various wave components and hence we often have to deal with convolutive signal mixtures, e.g., acoustic signals in reverberant indoor environments. In other words, assuming the propagation medium is linear, we have a matrix of convolutive propagation paths from the multiple sources to the multiple employed sensors. Since typically neither the original source signals nor the convolutive propagation paths are known *a priori*, we generally treat in this paper this overall array-based scene analysis problem as a *blind estimation problem*.

The special focus of this paper is on (1) solving the above-mentioned array signal processing problems using blind broadband adaptive MIMO (Multiple-Input and Multiple-Output) systems with explicit FIR models, and (2) to consider specifically the case of *time-varying MIMO mixing systems* within a *Bayesian framework*. In practice, time-varying mixing systems are very common, as they describe

- moving sources,
- moving reflectors/scattering objects,
- moving sensor array,

or a combination thereof. Accordingly, from an algorithmic point of view, this paper will bring together elements from Bayesian tracking (e.g., [1, 2]) and elements from blind adaptive signal processing (e.g., [3, 4, 5, 6]). Apart from the high-dimensionality of this tracking problem (on the order of the filter length L) compared to common 2D or 3D tracking problems, another important challenge from the signal-processing point of view is the blindness (well known tracking algorithms, such as Kalman filters or particle filters, are generally to be regarded as supervised algorithms). Moreover, an important objective in this paper is to make use of TRINICON ('TRiple-N ICA for CONvolute mixtures'), a previously introduced generic concept for broadband adaptive MIMO filtering, e.g., [4, 5, 6], using the technique of independent component analysis (ICA), e.g., [3], and in the *Bayesian TRINICON* framework, as presented in this paper, all the essential structures and degrees of freedom (see the references for more details) should be maintained.

The main three properties of TRINICON, making it an attractive vehicle for the present paper are: Its rigorous broadband formalism, its configurability to all the above tasks (blind, semi-blind, and also supervised [7, 8]), and its inherent exploitation of all possible statistical properties of the source signals ('TRiple-N'). Due to these properties, we can make use of the various relationships to most of the known blind, semi-blind, and supervised adaptive signal processing algorithms via the well defined degrees of freedom, and the fact that it has already led to various efficient time-domain and frequency-domain realizations, such as the blind unconstrained frequency-domain adaptive filter (now widely popular as *Independent Vector Analysis*) [6, 4].

Finally, an additional focus (3) of this paper is to formulate the (TRINICON-based) estimation and tracking on an arbitrary partly smooth manifold $\mathcal{M}$ in a way as shown previously in [9] for the non-Bayesian case. As shown in [9], being able to choose arbitrary partly smooth manifolds (a manifold is a topological space that is *locally* Euclidean [10, 11]) is a powerful approach to constrain the search space according to requirements of the desired application, or to further increase the convergence speed, or to reduce the computational complexity, or a combination thereof. Hence, it also allows us to

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extend the framework to the various semi-blind algorithms (including supervised algorithms as a special case), such as adaptive beamforming. Moreover, for blind adaptive filtering the *natural* manifold is of special importance (e.g., natural gradient descent [12]), and the general manifold-based approach in this paper will show how to incorporate this concept into the Bayesian tracking algorithms.

Generally, one of the main advantages of working with broadband MIMO FIR models is that we can, in principle, avoid any signal distortions in the source separation and signal enhancement tasks. Moreover, once we have obtained a robust estimate of the MIMO demixing system, we can infer most of the other relevant information for scene analysis, as mentioned above. A prominent example is acoustic source localization which is typically based on certain intermediate quantities, such as direction(s) of arrival (DOA) or time difference(s) of arrival (TDOA) in the classical approaches. It can be shown that these classical approaches can be related directly as special cases to the more 'holistic' approach outlined above (see Fig. 1), which explicitly takes into account the reverberant environment and the multiple, possibly simultaneously active sources (e.g., [4] and references therein). Similar relationships as in the forward adaptive

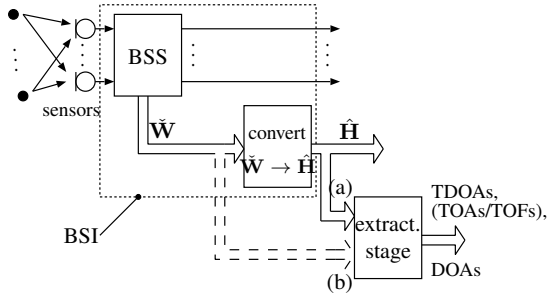


Fig. 1. Using blind source separation(BSS)/blind system identification (BSI) for source localization in convolutive environments.

filter problems (separation/interference cancellation, system identification) can also be established for inverse adaptive filter problems, such as blind dereverberation, e.g., [13].

Among the various known classes of Bayesian estimation and tracking algorithms (e.g., [1, 2]), we focus in this paper on Bayesian point estimators. This can be motivated by robustness and complexity considerations due to the typically very high dimensionality of the state space in our application, which is given by the filter length. Probably the best known and most popular classical tracking algorithm is the Kalman filter (KF) [14, 15], and its variants, such as the Extended Kalman filter (EKF) and the Unscented Kalman filter (UKF) [1, 2]. This class of algorithms can either be derived from a full Bayesian approach (under Gaussian assumptions, specified by mean and covariance) or *equivalently* as a point estimator of the mean vector (where the covariance estimate is obtained in the algorithm as a by-product). Coming from the field of adaptive array processing, and for the purpose of unifying these ideas with TRINICON, the latter approach is probably the more illustrative one in this paper. Indeed, the original paper by Thiele [14], referring directly to the earlier work by Gauss on the least-squares (LS) estimator, already describes this algorithm as a *dynamical model-based regularized LS problem* (directly reflected by the title of [14]), where the additive regularization term to the LS cost function is the *log prior* according to Bayes' rule. For adaptive signal processing in the special case of supervised adaptation, the (nonregularized) TRINICON-based adaptation was first contrasted with the Kalman-based (i.e., dynamically regularized RLS) adaptation in [16].

In this paper, we will confirm experimentally, that this additive regularization formulation also holds for the TRINICON optimization criterion, so that this idea of model-based regularization also carries over to the general blind and semi-blind broadband adaptive MIMO signal processing. A more detailed theoretical derivation of the Bayesian TRINICON optimization criterion can be found in [17]. Some early regularized ICA algorithms (not convolutive, not dynamically regularized, ICA only based on nongaussianity, not the full empirical risk minimization) can be found, e.g., in [18, 19]. Based on the Bayesian TRINICON optimization criterion, i.e., a suitably regularized version of the TRINICON criterion, and the manifold-based framework [9], the derivation of the generic algorithm in this paper follows from a *sequence estimation* in a similar way as shown in [20], by optimizing w.r.t. a *state vector* which collects the history (concatenation) of all sets of previous demixing filters.

2. GENERAL MIMO SETUP AND NOTATION

In this paper, we denote the original source signals by $s_q(n)$, $q = 1, \dots, Q$ and the captured sensor signals by $x_p(n)$, $p = 1, \dots, P$. We describe the MIMO mixing system by length- M FIR filters, where $h_{qp,\kappa}$, $\kappa = 0, \dots, M-1$ denote the coefficients of the FIR filter model from the q -th source signal $s_q(n)$ to the p -th sensor signal $x_p(n)$. We assume throughout this paper that $Q \leq P$. According to the optimization criterion, we are interested in finding a corresponding length- L FIR demixing system with coefficients $w_{pq,\kappa}$. This yields the output signals $y_q(n)$. As a compact formulation of the set of demixing filter coefficients and mixing filter coefficients we form the $PL \times Q$ demixing coefficient matrix

$$\tilde{\mathbf{W}} = \begin{bmatrix} \mathbf{w}_{11} & \cdots & \mathbf{w}_{1Q} \\ \vdots & \ddots & \vdots \\ \mathbf{w}_{P1} & \cdots & \mathbf{w}_{PQ} \end{bmatrix} \quad (1)$$

and the corresponding $QM \times P$ mixing coefficient matrix $\tilde{\mathbf{H}}$, respectively, where

$$\mathbf{h}_{qp} = [h_{qp,0}, \dots, h_{qp,M-1}]^T, \quad (2)$$

$$\mathbf{w}_{pq} = [w_{pq,0}, \dots, w_{pq,L-1}]^T \quad (3)$$

denote the coefficient vectors of the FIR subfilters of the MIMO systems, and superscript T denotes transposition of a vector or a matrix. The downwards pointing hat symbol on top of $\mathbf{W}$ in (1) serves to distinguish this *condensed* matrix from the corresponding larger matrix structure $\mathbf{W}$ as introduced below. The rigorous distinction between these different matrix structures is also an essential aspect of the general TRINICON framework, as shown later.

To model the time-varying nature of $\tilde{\mathbf{W}}$, we consider in this paper the following class of *state-space models* with *block time-index* m and a vector-valued function $\mathbf{g}(\cdot, \cdot)$:

$$\text{vec } \tilde{\mathbf{W}}(m+1) = \mathbf{g}(\text{vec } \tilde{\mathbf{W}}(m), \mathbf{u}(m)) + \mathbf{z}(m), \quad (4)$$

where $\mathbf{u}(m)$ denotes some possible innovation term, and $\mathbf{z}(m)$ denotes the process noise. As an important special case, this model includes the linear Gauss-Markov model with a (possibly time-varying) transition matrix $\mathbf{A}(m)$,

$$\text{vec } \tilde{\mathbf{W}}(m+1) = \mathbf{A}(m) \text{vec } \tilde{\mathbf{W}}(m) + \mathbf{u}(m) + \mathbf{z}(m), \quad (5)$$

where the process noise $\mathbf{z}(m)$ is assumed to be gaussian, described by a covariance matrix $\mathbf{Q}(m) = \mathcal{E}\{\mathbf{z}(m)\mathbf{z}^T(m)\}$.

3. TRINICON AS A GENERAL FRAMEWORK FOR BROADBAND SIGNAL PROCESSING ON CONVOLUTIVE MIXTURES AND EXTENSION TO THE BAYESIAN CASE

3.1. Optimization Criterion

In the following, to simplify the derivation, we work with an *expanded state vector* $\text{vec } \tilde{\boldsymbol{\theta}}(m)$ of size $(m+1)P^2L \times 1$ which captures the whole state evolution, i.e., the *sequence of coefficient vectors* $\text{vec } \tilde{\mathbf{W}}(i)$ from block time index $i = 0$ up to and including the current block time index $i = m$:

$$\begin{aligned} \text{vec } \tilde{\boldsymbol{\theta}}(m) &:= \left[\text{vec}^T \tilde{\mathbf{W}}(0), \dots, \text{vec}^T \tilde{\mathbf{W}}(m) \right]^T \\ &= \begin{bmatrix} \text{vec } \tilde{\boldsymbol{\theta}}(m-1) \\ \text{vec } \tilde{\mathbf{W}}(m) \end{bmatrix}. \end{aligned} \quad (6)$$

Taking into account the whole sequence $\tilde{\boldsymbol{\theta}}(m)$ for the optimization process rather than only the current coefficient matrix $\tilde{\mathbf{W}}(m)$, and using the dynamical model (4) in this way as a prior on $\tilde{\boldsymbol{\theta}}(m)$ are essentially the keys for obtaining the filter equations with tracking capability. Based on this state vector, we now formulate a criterion that is minimized with respect to the *sequence* of filter coefficient vectors up to block m , i.e., w.r.t. the expanded state vector $\text{vec } \tilde{\boldsymbol{\theta}}(m)$:

$$\begin{aligned} \mathcal{J}(m, \boldsymbol{\theta}(m)) &= \sum_{i=0}^m \beta(i, m) \tilde{\mathcal{J}}(i, \boldsymbol{\theta}) \\ &= \sum_{i=0}^m \beta(i, m) \left[\tilde{\mathcal{J}}_0(i, \boldsymbol{\theta}) + \tilde{\mathcal{J}}_R(i, \boldsymbol{\theta}) \right], \end{aligned}$$

where $\beta(i, m)$ is a weighting function defining different classes of algorithms [8] and allowing for online, offline, or block-online algorithms [6]. In this paper we choose

$$\beta(i, m) = \alpha \cdot \lambda^{m-i} \quad (7)$$

with $0 < \lambda \leq 1$ and an arbitrary real constant $\alpha > 0$.

The two terms $\tilde{\mathcal{J}}_0$ and $\tilde{\mathcal{J}}_R$ in

$$\tilde{\mathcal{J}}(m, \boldsymbol{\theta}(m)) = \tilde{\mathcal{J}}_0(m, \boldsymbol{\theta}(m)) + \tilde{\mathcal{J}}_R(m, \boldsymbol{\theta}(m)) \quad (8a)$$

denote the current data-based term and the regularization term, respectively, as discussed in the following subsections.

An important feature of the choice (7) for $\beta(i, m)$ is that it allows a recursive computation as

$$\mathcal{J}(m, \boldsymbol{\theta}(m)) = \lambda \mathcal{J}(m-1, \boldsymbol{\theta}(m-1)) + \alpha \tilde{\mathcal{J}}(m, \boldsymbol{\theta}(m)). \quad (8b)$$

3.1.1. Regularization Term

We first consider the regularization term $\tilde{\mathcal{J}}_R(m, \boldsymbol{\theta}(m))$. As it can be interpreted as a logarithmized prior on the convolutive demixing system using the state-space model, we write

$$\begin{aligned} \tilde{\mathcal{J}}_R(m, \boldsymbol{\theta}(m)) &= \\ &= f(\text{vec } \tilde{\mathbf{W}}(m) - \mathbf{g}(\text{vec } \tilde{\mathbf{W}}(m-1), \mathbf{u}(m-1))) \\ &= f(\text{vec } \tilde{\mathbf{W}}(m) - \mathbf{g}(\mathbf{W}_{P^2L \times mP^2L}^{01} \text{vec } \tilde{\boldsymbol{\theta}}(m-1), \mathbf{u}(m-1))) \end{aligned} \quad (9)$$

with a potential function $f(\cdot)$ and the window matrix $\mathbf{W}_{P^2L \times mP^2L}^{01} := [\mathbf{0}, \dots, \mathbf{0}, \mathbf{I}]$. The special case of the Gauss-Markov regularizer follows for

$$\mathbf{g}(\bullet, \bullet) = \mathbf{A}(m-1) \bullet + \bullet, \quad (10a)$$

$$f(\bullet) = \|\bullet\|_{\mathbf{Q}^{-1}(m-1)}^2 = \bullet^T \mathbf{Q}^{-1}(m-1) \bullet. \quad (10b)$$

For the recursive calculation of $\mathcal{J}$ we assume $\mathbf{A}(-1) = \mathbf{0}$ and $\mathbf{u}(-1) =: \mathbf{u}_0$ as initial values for $i = 0$.

3.1.2. Data Term

Various approaches exist to estimate the demixing matrix $\tilde{\mathbf{W}}$ by utilizing the fundamental source signal properties nongaussianity, nonwhiteness, and nonstationarity [3] which were all combined in TRINICON using the following data term [5]:

$$\begin{aligned} \tilde{\mathcal{J}}_0(m, \mathbf{W}(m)) &= \\ &= -\frac{1}{N} \sum_{j=mN_L}^{mN_L+N-1} \{\log(\hat{p}_{s,PD}(\mathbf{y}(j))) - \log(\hat{p}_{y,PD}(\mathbf{y}(j)))\} \end{aligned} \quad (11)$$

where $\hat{p}_{s,PD}(\cdot)$ and $\hat{p}_{y,PD}(\cdot)$ are assumed or estimated PD -variate source model (i.e., desired) pdf and output pdf, respectively. The index m denotes the block time index for a block of N output samples shifted by L samples relatively to the previous block.

To introduce an algorithm for broadband processing of convolutive mixtures, we formulate the convolution of the FIR demixing system of length L and with D time lags for each demixing filter output channel in the following matrix form [5]:

$$\mathbf{y}(n) = \mathbf{W}^T \mathbf{x}(n), \quad (12)$$

where n denotes the time index, and

$$\mathbf{x}(n) = [\mathbf{x}_1^T(n), \dots, \mathbf{x}_P^T(n)]^T, \quad (13)$$

$$\mathbf{y}(n) = [\mathbf{y}_1^T(n), \dots, \mathbf{y}_P^T(n)]^T, \quad (14)$$

$$\mathbf{x}_p(n) = [x_p(n), \dots, x_p(n-2L+1)]^T, \quad (15)$$

$$\mathbf{y}_q(n) = [y_q(n), \dots, y_q(n-D+1)]^T. \quad (16)$$

The parameter D in (16), $1 \leq D < L$, denotes the number of time lags taken into account to exploit the nonwhiteness of the source signals. $\mathbf{W}_{pq}$, $p = 1, \dots, P$, $q = 1, \dots, P$ denote $2L \times D$ *Sylvester matrices* that contain all coefficients of the respective filters in each column by successive shifting, i.e., the first column reads $[\mathbf{w}_{pq}^T, 0, \dots, 0]^T$, the second column $[0, \mathbf{w}_{pq}^T, 0, \dots, 0]^T$, etc. Finally, the $2PL \times PD$ matrix $\mathbf{W}$ combines all Sylvester matrices $\mathbf{W}_{pq}$.

The rigorous distinction between the different matrix structures $\tilde{\mathbf{W}}$ and $\mathbf{W}$ due to the nonwhiteness for $D > 1$ is also an essential aspect of the general TRINICON framework and leads to an important building block whose actual implementation is fundamental to the properties of the resulting algorithm, the so-called *Sylvester constraint (SC)* on the coefficient update, formally introduced in [6]. The Euclidean gradient of the data term can then be written as

$$\nabla_{\tilde{\mathbf{W}}} \tilde{\mathcal{J}}_0(m, \mathbf{W}) = \mathcal{SC} \left\{ \nabla_{\mathbf{W}} \tilde{\mathcal{J}}_0(m, \mathbf{W}) \right\}, \quad (17)$$

or using a fixed matrix $\mathbf{K}_{SC}$ (e.g., [9]),

$$\nabla_{\text{vec } \tilde{\mathbf{W}}} \tilde{\mathcal{J}}_0(m, \mathbf{W}) = \mathbf{K}_{SC} \nabla_{\text{vec } \mathbf{W}} \tilde{\mathcal{J}}_0(m, \mathbf{W}), \quad (18)$$

where

$$\nabla_{\mathbf{W}} \tilde{\mathcal{J}}_0(m, \mathbf{W}) = \sum_{j=iN_L}^{iN_L+N-1} \left[\mathbf{x}(j) \Phi_{s,PD}^T(\mathbf{y}(j)) - (\mathbf{W}^T)^+ \right], \quad (19)$$

with $^+$ denoting the pseudoinverse of a matrix, and with the generalized score function $\Phi_{s,PD}(\mathbf{y}(j))$ according to [13]. As an example, for convolutive blind source separation, the desired pdf is factorized w.r.t. the output channels. This leads to a concatenation of P individual score functions $\Phi_{q,D}$ of dimension D each, and for the special case of second-order statistics, i.e., multivariate Gaussian source models [6], they read $\Phi_{q,D}(\mathbf{y}_q(j)) = \mathbf{R}_{\mathbf{y}_q \mathbf{y}_q}^{-1}(i) \mathbf{y}_q(j)$.

3.2. Generic Newton-type algorithm for sequence estimation on arbitrary partly smooth manifolds for time-varying convolutive mixing systems

In [9] we presented a generic TRINICON-based Newton-type algorithm on arbitrary partly smooth manifolds, with multiple possible iterations per block. Since a manifold $\mathcal{M}$ is defined as a topological space that is locally Euclidean, local parameterizations in a Euclidean *tangent space* always exist. Hence, the idea in [9] was to perform each update step $\Delta \mathbf{T}(m)$ in the current tangent space, followed by the application of the local map $\mathbf{W}^\ell = \varphi_{\mathbf{W}^{\ell-1}}(\Delta \mathbf{T})$ back to the coefficient space.

In this paper, we apply this strategy for the more general case of *sequence estimation* by formulating the algorithm in terms of the corresponding *state vectors* for the coefficient matrices and the tangent space updates $\Delta \mathbf{T}(m)$ analogously to (6). This resulting algorithm can be simplified significantly in a number of steps (see [17] for detailed derivations), the most important steps resulting from picking only the last subvector of the new state vector, and by splitting the algorithm into a model-based *prediction part* and a data-based *correction part*. It turns out that this prediction-correction formulation is indeed made possible for arbitrary partly smooth manifolds by the chosen state-space model (4).

In the following pseudocode we summarize the generic Bayesian TRINICON algorithm on arbitrary partly smooth manifolds, based on the nonlinear state space model (4) with multiple iterations in Prediction-Correction formulation. It should be noted that this algorithm can be further developed into a Prediction-Innovation-Gain-Correction formulation analogously to the known classical Kalman-type tracking algorithms (see [17] for details).

$$\begin{aligned}
 & \text{Definitions:} \\
 \mathbf{F}(m-1) &= \nabla_{\mathbf{z}} \nabla_{\mathbf{z}}^T f(\mathbf{z}) \Big|_{\mathbf{z}=\text{vec } \tilde{\mathbf{W}}(m)-\mathbf{g}(\text{vec } \tilde{\mathbf{W}}(m-1), \mathbf{u}(m-1))} \\
 \mathbf{G}(m-1) &= \nabla_{\mathbf{z}} \mathbf{g}^T(\mathbf{z}, \mathbf{u}(m-1)) \Big|_{\mathbf{z}=\text{vec } \tilde{\mathbf{W}}(m-1)} \\
 \Xi' = \Xi(m-1, \mathbf{W}(m-1)) &= \frac{\partial \text{vec}^T(\varphi_{\mathbf{W}(m-1)})}{\partial \text{vec}(\Delta \mathbf{T}(m-1))} \Big|_{\Delta \mathbf{T}=0} \\
 & \text{Predictions:} \\
 \text{vec } \tilde{\mathbf{W}}^0(m) &= \mathbf{g}(\text{vec } \tilde{\mathbf{W}}(m-1), \mathbf{u}(m-1)) \\
 \tilde{\mathbf{P}}_{\Delta \mathbf{T}}(m | \mathbf{W}(m-1)) &= \\
 &= \frac{1}{\alpha} \mathbf{F}^{-1}(m-1) + \frac{1}{\lambda} \mathbf{G}(m-1) \mathbf{K}_{SC} \Xi'^T \\
 \tilde{\mathbf{P}}_{\Delta \mathbf{T}}(m-1 | \mathbf{W}(m-1)) \Xi' \mathbf{K}_{SC}^T \mathbf{G}^T(m-1) \\
 & \text{Corrections:} \\
 \text{For } \ell = 1, \dots, \ell_{\max}: \\
 \Xi = \Xi(m, \mathbf{W}^{\ell-1}(m)) &= \frac{\partial \text{vec}^T(\varphi_{\mathbf{W}^{\ell-1}(m)})}{\partial \text{vec}(\Delta \mathbf{T}(m))} \Big|_{\Delta \mathbf{T}(m)=0} \\
 \tilde{\mathbf{P}}_{\Delta \mathbf{T}}^{-1}(m | \mathbf{W}^{\ell-1}(m)) &= \Xi \mathbf{K}_{SC}^T \tilde{\mathbf{P}}_{\Delta \mathbf{T}}^{-1}(m | \mathbf{W}(m-1)) \mathbf{K}_{SC} \Xi^T \\
 &+ \alpha \Xi \nabla_{\text{vec } \mathbf{W}} \nabla_{\text{vec } \mathbf{W}}^T \tilde{\mathcal{J}}_0(m, \mathbf{W}^{\ell-1}(m)) \Xi^T \\
 \text{vec } \Delta \mathbf{T}^\ell(m) &= -\mu \tilde{\mathbf{P}}_{\Delta \mathbf{T}}(m | \mathbf{W}^{\ell-1}(m)) \sum_{i=0}^{\infty} \beta(i, m) \Xi \\
 &\cdot \text{vec} \left\{ \nabla_{\mathbf{W}(m)} \tilde{\mathcal{J}}_0(m, \mathbf{W}(m)) \right\} \Big|_{\mathbf{W}=\mathbf{W}^{\ell-1}(m)} \\
 \tilde{\mathbf{W}}^\ell(m) &= \mathcal{S} \left\{ \varphi_{\mathbf{W}^{\ell-1}(m)}(\Delta \mathbf{T}^\ell(m)) \right\} \\
 \text{endfor} \\
 \tilde{\mathbf{W}}(m) &= \tilde{\mathbf{W}}^{\ell_{\max}}(m) \\
 \tilde{\mathbf{P}}_{\Delta \mathbf{T}}(m | \mathbf{W}(m)) &= \tilde{\mathbf{P}}_{\Delta \mathbf{T}}(m | \mathbf{W}^{\ell_{\max}-1}(m))
 \end{aligned}$$

4. APPLICATION TO BROADBAND BSS FOR TIME-VARYING CONVOLUTIVE MIXING SYSTEMS

To evaluate the effectiveness of the Bayesian TRINICON approach with state-space model, we consider blind source separation of speech signals in a time-varying reverberant environment. Our measurements were performed in a real office environment (with curtains and carpet). The reverberation time T_{60} was approximately 300ms and the sampling rate was 16kHz. We placed two (spatially fixed) speech sources (using loudspeakers) and two (spatially fixed) microphones in this room. To simulate the time-variance of the 2×2 acoustic mixing system in a reproducible way, we placed a rotating panel ($\approx 2m^2$) on a stepper motor nearby the microphones. Using the stepper motor, this panel was rotated in 0.2° -steps so that in total, we obtained a set of 900 measured 2×2 MIMO systems for the entire rotation by 180° .

We evaluate the BSS separation performance in terms of the signal-to-interference ratio (SIR) improvements on the outputs of the demixing system, or, more precisely, the arithmetic average between the SIR improvements in the two output signals $y_1(n)$ and $y_2(n)$. Note that the precise SIR improvements can be calculated only in simulations when the measured *mixing* systems are available as a reference. Hence, in the case of time-varying mixing systems we used our entire set of 900 measured MIMO systems as described above in order to calculate the SIR improvements, i.e., to show the tracking performance during the online BSS operation.

Figure 2 shows the SIR improvements obtained by two different algorithms in the time-varying environment after reaching their steady-state performance (i.e., after their initial convergence). The

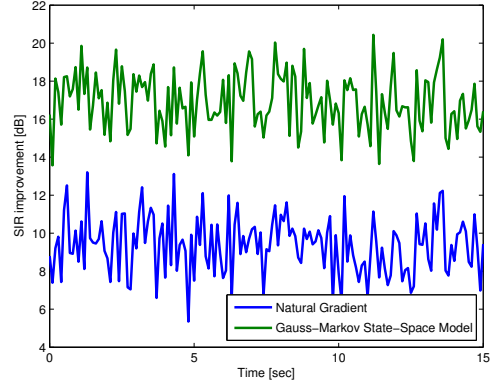


Fig. 2. Simulation results for blind source separation in a continuously time-varying environment.

lower (blue) SIR improvement curve shows the performance of the natural gradient-based broadband BSS algorithm (with a fixed Gaussian prior for the demixing system). We regard this as a state-of-the-art baseline algorithm (e.g., [4, 5, 6]). The upper (green) SIR improvement curve was obtained using a corresponding broadband BSS algorithm with a Gauss-Markov state-space model leading to the novel blind Kalman-like BSS algorithm (also on the natural manifold, $\varphi_{\mathbf{W}^{\ell-1}}(\Delta \mathbf{T}) = \mathbf{W}^{\ell-1} + \mathbf{W}^{\ell-1} \Delta \mathbf{T}$, see [9]). In this case, the Kalman gain computation (including the inverse Hessian) was performed in the frequency domain for computational efficiency.

5. CONCLUSIONS

In this paper we have presented a unified Bayesian adaptive signal processing and tracking framework. The resulting SIR performance for the BSS example in time-varying environments shows a clear improvement over the baseline algorithm.

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