

THE CHORD GAP DIVERGENCE AND A GENERALIZATION OF THE BHATTACHARYYA DISTANCE

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ABSTRACT

We introduce a novel family of distances, called the chord gap divergences, that generalizes the Jensen/Burbea-Rao distances and study its properties. It follows a generalization of the statistical Bhattacharyya distance that is frequently met in applications. We then report an iterative concave-convex procedure for computing centroids, and analyze the performance of the k -means++ clustering with respect to that new dissimilarity measure by introducing the Taylor-Lagrange remainder form of skew Jensen divergences.

Index Terms— Jensen/Burbea-Rao divergence, Bregman divergence, Jensen-Bregman divergence, centroid, k -means++

1. INTRODUCTION

In many applications, one faces the crucial dilemma of choosing an appropriate distance $D(\cdot, \cdot)$ between data. In some cases, those distances can be picked up *a priori* from well-grounded principles (e.g., Kullback-Leibler distance in statistical estimation [1]). In other cases, one is rather left at testing several distances [2], and choose *a posteriori* the distance that yielded the best performance. For the latter cases, it is judicious to consider a family of parametric distances $D_\alpha(\cdot, \cdot)$, and learn [3] the hyperparameter α according to the application at hand and potentially the dataset (distance selection). Thus it is interesting to consider parametric generalizations of common distances [4] to improve performance in applications.

Some distances can be designed from inequality gaps [5, 6]. For example, the Jensen divergence $J_F(p, q)$ (also called the Burbea-Rao divergence [5, 7]) is designed from the inequality gap of Jensen inequality $F\left(\frac{p+q}{2}\right) \leq \frac{F(p)+F(q)}{2}$ that holds for a strictly convex function F : $J_F(p, q) = \frac{F(p)+F(q)}{2} - F\left(\frac{p+q}{2}\right)$. We can extend the Jensen divergence to a parametric family of skew Jensen divergences J_F^α (with $\alpha \in (0, 1)$) built on the convex inequality gap $F((1-\alpha)p + \alpha q) \leq (1-\alpha)F(p) + \alpha F(q)$:

$$J_F^\alpha(p : q) = (1-\alpha)F(p) + \alpha F(q) - F((1-\alpha)p + \alpha q), \quad (1)$$

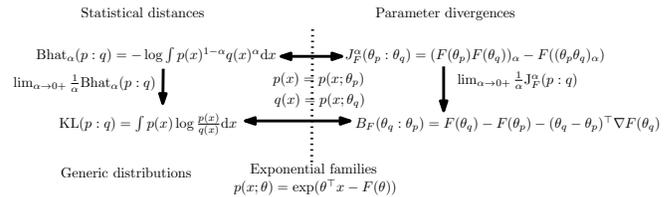


Fig. 1. Links between the statistical skew Bhattacharyya distances and parametric skew Jensen divergences when distributions belong to the same exponential family.

satisfying $J_F^\alpha(q : p) = J_F^{1-\alpha}(p : q)$ and $J_F^{\frac{1}{2}}(p : q) = J_F(p, q)$. Here the ‘:’ notation emphasizes the fact that the distance is potentially asymmetric $J_F^\alpha(p : q) \neq J_F^\alpha(q : p)$. The term divergence is used in information geometry [9] to refer to the smoothness of the distance that yields an information-geometric structure of the space induced by the divergence. Let $[p, q] = \{(pq)_\lambda := (1-\lambda)p + \lambda q, \lambda \in [0, 1]\}$ denote the line segment with endpoints p and q . Then we can rewrite [8] Eq.1:

$$J_F^\alpha(p : q) = (F(p)F(q))_\alpha - F((pq)_\alpha). \quad (2)$$

In applications, it is rather the relative comparisons of distances rather than their absolute values that is important. Thus we may multiply a distance by any positive scaling factor and include it in the class of that distance. When F is strictly convex and differentiable, the class of Jensen divergences include in the limit cases the Bregman divergences [7, 10, 11]: $\lim_{\alpha \rightarrow 0^+} \frac{J_F^\alpha(p : q)}{\alpha} = B_F(q : p)$, and $\lim_{\alpha \rightarrow 1^-} \frac{J_F^\alpha(p : q)}{1-\alpha} = B_F(p : q)$, where

$$B_F(p : q) = F(p) - F(q) - (p - q)^\top \nabla F(q), \quad (3)$$

is the Bregman divergence [13]. Overall, one may define the smooth parametric family of scaled skew Jensen divergences: $sJ_F^\alpha(p : q) = \frac{1}{\alpha(1-\alpha)} J_F^\alpha(p : q)$ that encompasses the Bregman divergence $B_F(p : q)$ and the reverse Bregman divergence $B_F(q : p)$ in limit cases (with $\alpha \in \mathbb{R}$).

There is a nice relationship between Jensen divergences operating on parameters (e.g., vectors, matrices) and a class

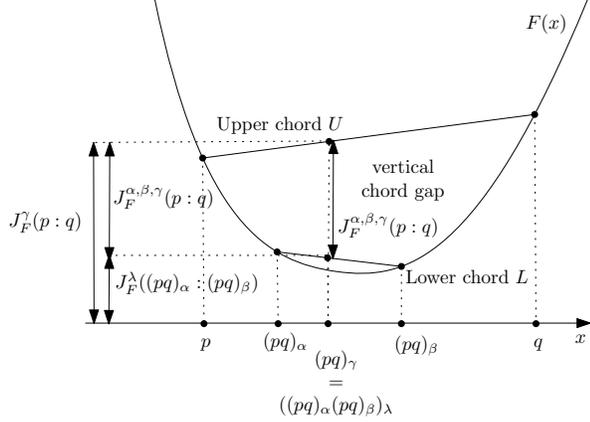


Fig. 2. The triparametric chord gap divergence.

of statistical distances between probability distributions (Figure 1): Let $\{p(x; \theta)\}_\theta$ be an exponential family [11] (includes the Gaussian family and the finite discrete “multinoulli” family) with convex cumulant function $F(\theta)$. Then the skew Bhattacharyya distance [14]

$$\text{Bhat}_\alpha(p : q) = -\log \int p(x)^{1-\alpha} q(x)^\alpha dx, \quad (4)$$

between two distributions belonging to the same exponential family amounts to a skew Jensen divergence [7]:

$$\text{Bhat}(p(x; \theta_1) : p(x; \theta_2)) = J_F^\alpha(\theta_1 : \theta_2). \quad (5)$$

We further check that $\lim_{\alpha \rightarrow 0^+} \frac{1}{\alpha} \text{Bhat}_\alpha(p : q) = \text{KL}(p : q)$ and $\lim_{\alpha \rightarrow 1^-} \frac{1}{1-\alpha} \text{Bhat}_\alpha(p : q) = \text{KL}(q : p)$ where $\text{KL}(p : q) = \int p(x) \log \frac{p(x)}{q(x)} dx$ is the Kullback-Leibler divergence.

In statistical signal processing, information fusion and machine learning, one often considers the skew Bhattacharyya distance [15, 16, 17] or the Chernoff distance [18, 19, 20] for exponential families (e.g., Gaussian/multinoulli): This highlights the important role in disguise of the equivalent skew Jensen divergences (see Eq. 5).

The paper is organized as follows: Section 2 introduces the novel triparametric family of chord gap divergences that generalizes the skew Jensen divergences (§2.1), describes several properties (§2.2), and deduces a generalization of the statistical Bhattacharyya distance (§2.3). Section 3 considers the calculation of the centroid (§3.1) for the chord gap divergences, and report probabilistic guarantee of the k -means++ seeding (§3.2) by highlighting the Taylor-Lagrange forms of those divergences.

2. THE CHORD GAP DIVERGENCE

2.1. Definition

Let $F : \mathcal{X} \rightarrow \mathbb{R}$ be a *strictly convex* function. For $\alpha, \beta \in (0, 1)$ with $\alpha \neq \beta$, the chord

$L = [((pq)_\alpha, F((pq)_\alpha)), ((pq)_\beta, F((pq)_\beta))]$ is *below* the distinct chord $U = [(p, F(p)), (q, F(q))]$. Thus we can define a divergence [21] as the vertical gap between these two chords for a given coordinate $x \in [(pq)_\alpha, (pq)_\beta]$:

$$J_F^{\alpha, \beta, \gamma}(p : q) = (F(p)F(q))_\gamma - (F((pq)_\alpha)F((pq)_\beta))_\lambda, \quad (6)$$

such that $((pq)_\alpha, (pq)_\beta)_\lambda = (pq)_\gamma$ with $\gamma \in (\alpha, \beta)$ (Figure 2). A calculation shows that $\lambda = \lambda(\alpha, \beta, \gamma) = \frac{\gamma - \alpha}{\beta - \alpha}$ or $\gamma = \lambda(\beta - \alpha) + \alpha$ for $\lambda \in [0, 1]$ when $\alpha \neq \beta$, so that we get $J_F^{\alpha, \beta, \gamma}(p : q) = (F(p)F(q))_\gamma - (F((pq)_\alpha)F((pq)_\beta))_{\frac{\gamma - \alpha}{\beta - \alpha}}$.

2.2. Properties of the chord gap divergence

We have $J_F^{\alpha, \alpha, \alpha}(p : q) = J_F^\alpha(p : q)$, $J_F^{0, 1, \gamma}(p : q) = J_F^\gamma(p : q)$ and $J_F^{\alpha, \beta, \gamma}(q : p) = J_F^{1-\alpha, 1-\beta, 1-\gamma}(p : q)$ since $\lambda(1-\alpha, 1-\beta, 1-\gamma) = \frac{\gamma - \alpha}{\beta - \alpha} = \lambda(\alpha, \beta, \gamma)$ using the fact that $(ab)_{1-\delta} = (ba)_\delta$ for $\delta \in [0, 1]$. Thus we have $J_F^{1-\alpha, 1-\alpha, 1-\alpha}(p : q) = J_F^\alpha(q : p)$. For $\beta = 1 - \alpha$, $\gamma = \frac{1}{2}$ (and $\lambda = \frac{1}{2}$), the chord gap divergence amounts to a scaled symmetrized skew Jensen divergences [12] (Eq. 35).

We can express the chord gap divergence as the difference of two skew Jensen divergences (Figure 2):

$$J_F^{\alpha, \beta, \gamma}(p : q) = J_F^\gamma(p : q) - J_F^\lambda((pq)_\alpha : (pq)_\beta), \quad (7)$$

with $\lambda = \frac{\gamma - \alpha}{\beta - \alpha}$ or $\gamma = \lambda(\beta - \alpha) + \alpha$ for $\lambda \in [0, 1]$ and $\gamma \in [\alpha, \beta]$. Thus the chord gap divergence can be interpreted as a *truncated* skew Jensen divergence.

A biparametric subfamily $J_F^{\beta, \gamma}$ of $J_F^{\alpha, \beta, \gamma}$ is obtained by setting $\alpha = 0$ so that $(pq)_\alpha = p$, so that the two upper/lower chords L and U coincide at extremity p :

$$\begin{aligned} J_F^{\beta, \gamma}(p : q) &= (F(p)F(q))_\gamma - (F(p)F((pq)_\beta))_{\frac{\gamma}{\beta}}, \quad (8) \\ &= \gamma \left(\left(\frac{1}{\beta} - 1 \right) F(p) + F(q) - \frac{1}{\beta} F((pq)_\beta) \right). \end{aligned}$$

When $\beta = \frac{1}{2}$, we find that $J_F^{\frac{1}{2}, \gamma}(p : q) = 2\gamma J_F(p : q)$, the ordinary (γ -scaled) Jensen divergence. When $\beta \rightarrow 0$, we have $\lim_{\beta \rightarrow 0} \frac{1}{\gamma} J_F^{\beta, \gamma}(p : q) = B_F(q : p)$ (with $\gamma \in (0, \beta)$) since $-\frac{1}{\beta} F((pq)_\beta) \simeq -\frac{1}{\beta} - (q - p)^\top \nabla F(p)$ using a first-order Taylor expansion.

Matrix chord gap divergences can be obtained by taking strictly convex matrix generators [13] (e.g., $F(X) = -\log \det |X|$) for symmetric positive definite matrices $X \in \mathbb{P}_{++}$, $\mathbb{P}_{++} = \{X : X \succ 0\}$ denote the space of positive definite matrices, a convex cone. This may be useful in applications based on covariance matrices [13].

2.3. Generalized Bhattacharyya distances

The interpretation given in Eq. 7 yields a triparametric family of Bhattacharyya statistical distances [14] between members $p(x) = p(x; \theta_p)$ and $q(x) = p(x; \theta_q)$ of the *same* exponential family: $\text{Bhat}_{\alpha, \beta, \gamma}(\theta_p : \theta_q) = \text{Bhat}_\gamma(\theta_p : \theta_q) -$

Bhat $_{\lambda}((\theta_p\theta_q)_{\alpha} : (\theta_p\theta_q)_{\beta})$. It follows that

$$\text{Bhat}_{\alpha,\beta,\gamma}(\theta_p : \theta_q) = -\log \frac{\int p(x; \theta_p)^{1-\gamma} p(x; \theta_q)^{\gamma} dx}{\int p(x; (\theta_p\theta_q)_{\alpha})^{1-\lambda} p(x; (\theta_p\theta_q)_{\beta})^{\lambda} dx}. \quad (9)$$

Note that when $\alpha = \beta$, we have $p(x; (\theta_p\theta_q)_{\alpha})^{1-\lambda} p(x; (\theta_p\theta_q)_{\beta})^{\lambda} = p(x; (\theta_p\theta_q)_{\alpha})$ and therefore the denominator is $\int p(x; (\theta_p\theta_q)_{\alpha}) dx = 1$, and we recover the skew Bhattacharyya distance, as expected.

We shall extend the generalized Bhattacharyya divergence of Eq. 9 to arbitrary distributions by generalizing the notion of interpolated distribution $p(x; (\theta_p\theta_q)_{\delta}) = \Gamma_{\delta}(p(x; \theta_p), p(x; \theta_q))$. When δ ranges from 0 to 1, we obtain a Bhattacharyya arc linking $p(x; \theta_p)$ to $p(x; \theta_q)$ (exponential or e -geodesic in information geometry [9]). We define:

$$\Gamma_{\delta}(p(x), q(x)) = \frac{p(x)^{1-\delta} q(x)^{\delta}}{Z_{\delta}(p(x) : q(x))},$$

with $Z_{\delta}(p(x) : q(x)) = \int p(x)^{1-\delta} q(x)^{\delta} d\nu(x)$. Note that we need the integral to converge properly in order to define $\Gamma_{\delta}(p(x), q(x))$. This always holds for distributions belonging to the same exponential families since $(\theta_p\theta_q)_{\delta}$ is guaranteed to belong to the natural parameter space, and

$$Z_{\delta}(p(x; \theta_p) : p(x; \theta_q)) = \exp(-J_F^{\delta}(\theta_p : \theta_q)). \quad (10)$$

By extension, the triparametric Bhattacharyya distance can be defined by:

$$\text{Bhat}^{\alpha,\beta,\gamma}(p(x) : q(x)) = -\log \left(\frac{\int p(x)^{1-\gamma} q(x)^{\gamma} d\nu(x)}{\Gamma_{\alpha}(p(x), q(x))^{1-\gamma} \Gamma_{\beta}(p(x), q(x))^{\gamma}} \right). \quad (11)$$

Thus we explicitly define the generalized Bhattacharyya distance by:

$$\text{Bhat}^{\alpha,\beta,\gamma}(p(x) : q(x)) = -\log \left(\frac{\int p(x)^{1-\gamma} q(x)^{\gamma} d\nu(x)}{\left(\int \frac{p(x)^{1-\alpha} q(x)^{\alpha} d\nu(x)}{\int p(x)^{1-\alpha} q(x)^{\alpha} d\nu(x)} \right)^{1-\lambda} \left(\int \frac{p(x)^{1-\beta} q(x)^{\beta} d\nu(x)}{\int p(x)^{1-\beta} q(x)^{\beta} d\nu(x)} \right)^{\lambda}} \right)$$

Notice that when $\alpha = \beta$, for any $\lambda \in [0, 1]$, the denominator collapses to one, and we find that $\text{Bhat}_{\alpha,\beta,\gamma}(p(x) : q(x)) = \text{Bhat}_{\alpha}(p(x) : q(x))$, as expected.

For multivariate Gaussians/normals belonging to the family $\{\mathcal{N}(\mu, \Sigma) : \mu \in \mathbb{R}^d, \Sigma \in \mathbb{P}_{++}^d\}$, we have the natural parameter [22] $\theta = (v, M) = (\Sigma^{-1}\mu, -\frac{1}{2}\Sigma^{-1})$, and the cumulant function $F(v, M) = \frac{d}{2} \log 2\pi - \frac{1}{2} \log | -2M | - \frac{1}{4} v^{\top} M^{-1} v$ that can also be expressed in the usual parameters $F(\mu, \Sigma) = \frac{1}{2} \log(2\pi)^d |\Sigma| + \frac{1}{2} \mu^{\top} \Sigma^{-1} \mu$. We have $(\theta_p\theta_q)_{\delta} = ((1-\delta)\Sigma_p^{-1}\mu_p + \delta\Sigma_q^{-1}\mu_q, -\frac{1-\delta}{2}\Sigma_p^{-1} - \frac{\delta}{2}\Sigma_q^{-1})$ so that we get [23]: $J_F^{\alpha}(p(x; \mu_p, \Sigma_p) : p(x; \mu_q, \Sigma_q)) = \frac{\alpha(1-\alpha)}{2} \Delta\mu^{\top} ((1-\alpha)\Sigma_p + \alpha\Sigma_q)^{-1} \Delta\mu + \frac{1}{2} \log \frac{|(1-\alpha)\Sigma_p + \alpha\Sigma_q|}{|\Sigma_p|^{1-\alpha} |\Sigma_q|^{\alpha}}$ with $|\cdot|$ the determinant and $\Delta\mu = \mu_q - \mu_p$. This gives a closed-form formula for $\text{Bhat}^{\alpha,\beta,\gamma}$ for multivariate Gaussians.

3. CENTROID-BASED CLUSTERING

Bhattacharyya clustering is often used in statistical signal processing, information fusion, and machine learning (see [15, 24, 25, 26] for some illustrative examples). Popular clustering algorithms are center-based clustering, where each cluster stores a prototype (a representative element), and each datum is assigned to the cluster with the closest prototype wrt. a distance function. The cluster prototypes are then updated, and the algorithm iterates until (local) convergence. This scheme includes the k -means and the k -medians [27]. Lloyd k -means heuristic updates the prototype c of a cluster X by choosing its center of mass $c = \frac{1}{|X|} \sum_{x \in X} x$ that minimizes the cluster variance: $\min_c \sum_{x \in X} \|x - c\|^2$ (this holds for any Bregman divergence too [11]).

3.1. Chord gap divergence centroid

We extend k -means for a weighted point set

$$\mathcal{P} = \{(w_1, p_2), \dots, (w_n, p_n)\},$$

with $w_i > 0$ and $\sum_i w_i = 1$, using the chord gap divergence by solving the following minimization problem: $\min_x E(x) = \sum_{i=1}^n w_i J_F^{\alpha,\beta,\gamma}(p_i : x)$. By expanding the chord gap divergence formula and removing all terms independent of x , we obtain an equivalent minimization problem as a difference of convex function programming [28]: $\min_x E(x) = \min_x A(x) - B(x)$ with $A(x) = \sum_{i=1}^n (F(p_i)F(x))_{\gamma}$ and $B(x) = \sum_{i=1}^n (F((p_i x)_{\alpha})F((p_i x)_{\beta}))_{\lambda}$, both strictly convex functions. It follows a concave-convex procedure [29] (CCCP) solving locally $\min_x A(x) - B(x)$: initialize $x_0 = p_1$ and then iteratively update $\nabla A(x_{t+1}) = \nabla B(x_t)$. When the reciprocal gradient ∇A^{-1} is available in closed form, we end up with $x_{t+1} = \nabla A^{-1}(\nabla B(x_t))$. Since we have $\nabla A(x) = n\gamma \nabla F(x)$ and $\nabla B(x) = \sum_i (1-\lambda)\alpha \nabla F((p_i x)_{\alpha}) + \lambda\beta \nabla F((p_i x)_{\beta})$, the update rule is

$$x_{t+1} = \nabla F^{-1} \left(\frac{1}{\gamma} \sum_i w_i ((1-\lambda)\alpha \nabla F((p_i x_t)_{\alpha}) + \lambda\beta \nabla F((p_i x_t)_{\beta})) \right).$$

When $\alpha = \beta = \gamma$, we find the simplified update rule $x_{t+1} = \nabla F^{-1}(\sum_i w_i \nabla F((p_i x_t)_{\alpha}))$ corresponding to the skew Jensen divergences [7]. Note that it is enough to improve iteratively the prototypes to get a variational Lloyd's k -means [30].

3.2. Performance analysis of k -means++

For high-performance clustering, one may use k -means++ [31] that is a guaranteed probabilistic initialization of the cluster prototypes. To get an expected competitive ratio [31] of $2U^2(1+V)(2+\log k)$ [30], we need to upper bound: (i) U such that the divergence $D = J_F^{\alpha,\beta,\gamma}$ satisfies the U -triangular inequality $D(x : z) \leq U(D(x : y) + D(y : z))$,

and (ii) V such that the divergence satisfies the symmetric inequality $D(y : x) \leq VD(x : y)$. The proof follows the proof reported in [30] for total Jensen divergences once we can express the divergences in their Taylor-Lagrange forms $D(p : q) = (p - q)^\top H_D(p : q)(p - q)$ where $H_D(p : q) \succ 0$. For example, the Taylor-Lagrange form of the Bregman divergence [32] is obtained from a first-order Taylor expansion with the exact Lagrange remainder:

$$B_F(p : q) = \frac{1}{2}(p - q)^\top \nabla^2 F(\xi)(p - q), \quad (12)$$

for some $\xi \in [p, q]$. This expression can be interpreted as a squared Mahalanobis distance $M_Q(p, q) = (p - q)^\top Q(p - q)$ with precision matrix $Q = \frac{1}{2}\nabla^2 F(\xi)$ depending on p and q . Any squared Mahalanobis distance satisfies $U = 2$ (see [33]) and $V = 1$, and can be interpreted as a squared norm-induced distance: $M_Q(p, q) = \|Q^{\frac{1}{2}}(p - q)\|_2^2$.

We report the Taylor-Lagrange form of the skew Jensen divergences: There exists $\xi_1, \xi_2 \in [p, q]$, such that the skew Jensen divergence can be expressed as $J_F^\alpha(p : q) = (p - q)^\top H_F^\alpha(p : q)(p - q)$, with

$$H_F^\alpha(p : q) = \frac{1}{2}\alpha(1 - \alpha)(\alpha\nabla^2 F(\xi_1) + (1 - \alpha)\nabla^2 F(\xi_2)). \quad (13)$$

The proof relies on introducing the skew Jensen-Bregman (JB) divergence [7] defined by

$$JB_F^\alpha(p : q) = (1 - \alpha)B_F(p : (pq)_\alpha) + \alpha B_F(q : (pq)_\alpha), \quad (14)$$

and observing the $JB_F^\alpha(p : q) = J_F^\alpha(p : q)$ since $p - (pq)_\alpha = \alpha(p - q)$ and $q - (pq)_\alpha = (1 - \alpha)(q - p)$ (and therefore the $\nabla F((pq)_\alpha)$ -terms cancel out). Then we apply the Taylor-Lagrange form of Bregman divergences of Eq. 12 to get the result. Notice that when $\alpha \rightarrow 0$ or $\alpha \rightarrow 1$, the scaled skew Jensen difference tend to Bregman divergences, and we have $\lim_{\alpha \rightarrow 1} \frac{J_F^\alpha(p : q)}{\alpha(1 - \alpha)} = \frac{1}{2}(p - q)^\top \nabla^2 F(\xi_1)(p - q) = B_F(p : q)$ for $\xi_1 \in [p, q]$, and $\lim_{\alpha \rightarrow 0} \frac{J_F^\alpha(p : q)}{\alpha(1 - \alpha)} = \frac{1}{2}(p - q)^\top \nabla^2 F(\xi_2)(p - q) = B_F(q : p)$ for $\xi_2 \in [p, q]$, as expected.

Using expression of Eq. 7 for the chord gap divergence, and the fact that $(pq)_\alpha - (pq)_\beta = (\alpha - \beta)(q - p)$, we get the Taylor-Lagrange form of the chord gap divergence $J_F^{\alpha, \beta, \gamma} = (p - q)^\top H_F^{\alpha, \beta, \gamma}(p : q)(p - q)$ with

$$\begin{aligned} H_F^{\alpha, \beta, \gamma}(p : q) &= \frac{1}{2}\gamma(1 - \gamma)\nabla^2 F(\xi') - \frac{1}{2}\lambda(1 - \lambda)(\alpha - \beta)^2\nabla^2 F(\xi''), \quad (15) \\ &= \frac{1}{2}(\gamma(1 - \gamma)\nabla^2 F(\xi') - (\gamma - \alpha)(\gamma - \beta)\nabla^2 F(\xi'')), \quad (16) \end{aligned}$$

for $\xi', \xi'' \in \mathcal{X}$.

When dealing with a finite (weighted) point set \mathcal{P} , let

$$\rho = \frac{\sup_{\xi', \xi'', p, q \in \text{co}(\mathcal{P})} \|(\nabla^2 F(\xi'))^{\frac{1}{2}}(p - q)\|}{\inf_{\xi', \xi'', p, q \in \text{co}(\mathcal{P})} \|(\nabla^2 F(\xi''))^{\frac{1}{2}}(p - q)\|} < \infty, \quad (17)$$

where $\text{co}(\mathcal{P})$ denotes the convex closure of \mathcal{P} . Then it comes that $U = O_\rho(1)$ and $V = O_\rho(1)$ so that k -means++ probabilistic seeding is $\bar{O}_\rho(\log k)$ competitive for the chord gap divergence.

4. CONCLUDING REMARKS

We introduced the chord gap divergence as a generalization of the skew Jensen divergences [7, 13], studied its properties and obtained a generalization of the skew Bhattacharyya divergences. We showed that the chord gap divergence centroid can be obtained using a convex-concave iterative procedure [7], and analyzed the k -means++ [31] performance by giving the Taylor-Lagrange forms of the skew Jensen and chord gap divergences. We expect our contributions to be useful for the signal processing, information fusion and machine learning communities where the Bhattacharyya [34, 35] or Chernoff information [2, 18] is often used. In practice, the triparametric chord gap divergence shall be tuned according to the application at hand (and the dataset for supervised tasks using cross-validation for example). JavaTM source code is available for reproducible research: www.lix.polytechnique.fr/~nielsen/CGD/

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